

Elements of Machine Learning & Data Science

Winter semester 2023/24

Lecture 4 – Probability Density Estimation I 20.10.2023

Prof. Bastian Leibe

Machine Learning Topics

- 1. Introduction to ML
- **2. Probability Density Estimation**
- 3. Linear Discriminants
- 4. Linear Regression
- 5. Logistic Regression
- 6. Support Vector Machines
- 7. AdaBoost
- 8. Neural Network Basics Mixtures of Gaussians

Parametric Methods & ML-Algorithm

Nonparametric Methods

& EM-Algorithm

Bayes Classifiers

Recap: The Gaussian (or Normal) Distribution

• One-dimensional (univariate) case:

$$
\mathcal{N}(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)
$$

Mean Variance

• Multi-dimensional (multivariate) case:

$$
\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{D/2} |\boldsymbol{\Sigma}|^{1/2}} \exp\left(-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^{\mathsf{T}} \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right)
$$

Mean Covariance
vector matrix

Probability Density Estimation

- 1. Probability Distributions
- **2. Parametric Methods**
- 3. Nonparametric Methods
- 4. Mixture Models
- 5. Bayes Classifier
- 6. K-NN Classifier

Parametric Methods

- In parametric methods, we assume that we know the parametric form of the underlying data distribution.
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Parametric Methods

- In parametric methods, we assume that we know the parametric form of the underlying data distribution.
	- I.e., the equation of the pdf with parameters θ .
- Goal: Estimate θ from training data $\mathcal{X} = {\mathbf{x}_1, \dots, \mathbf{x}_N}$.
- Likelihood of θ :

 $L(\theta) = p(\mathcal{X}|\theta)$

Probability that the data X *was indeed generated by a distribution with parameters* θ *.* Example: $p(x) = \mathcal{N}(x|\mu, \sigma)$

Maximum Likelihood Approach

- Idea: Find optimal parameters by maximizing $L(\theta)$.
- Computation of the likelihood:
	- Single data point (e.g., for Gaussian):

$$
p(x_n|\theta) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(x_n - \mu)^2}{2\sigma^2}\right)
$$

• Assumption: all data points are independent

$$
L(\theta) = p(\mathcal{X}|\theta) = \prod_{n=1}^{N} p(x_n|\theta)
$$

• Negative Log-Likelihood ("Energy"):

$$
E(\theta) = -\ln L(\theta) = -\sum_{n=1}^{N} \ln p(x_n|\theta)
$$

Maximizing the likelihood minimizing the negative log-likelihood.

- Minimizing the negative log-likelihood:
	- Take the derivative and set it to zero.

$$
\frac{\partial}{\partial \theta} E(\theta) = -\frac{\partial}{\partial \theta} \sum_{n=1}^{N} \ln p(x_n|\theta) = -\sum_{n=1}^{N} \frac{\frac{\partial}{\partial \theta} p(x_n|\theta)}{p(x_n|\theta)} \stackrel{!}{=} 0
$$

• Log-likelihood for Normal distribution (1D case):

$$
\frac{\partial}{\partial \mu} E(\mu, \sigma) = -\sum_{n=1}^{N} \frac{\frac{\partial}{\partial \mu} p(x_n | \mu, \sigma)}{p(x_n | \mu, \sigma)}
$$
\n
$$
= -\sum_{n=1}^{N} -\frac{\lambda (x_n - \mu)}{\lambda \sigma^2} \frac{p(x_n | \mu, \sigma)}{p(x_n | \mu, \sigma)}
$$
\n
$$
\frac{\partial}{\partial \mu} p(x_n | \mu, \sigma) = -\frac{2(x_n - \mu)}{2\sigma^2} p(x_n | \mu, \sigma)
$$

$$
= \frac{1}{\sigma^2} \sum_{n=1}^{N} (x_n - \mu) = \frac{1}{\sigma^2} \left(\sum_{n=1}^{N} x_n - N\mu \right) \stackrel{!}{=} 0 \iff \hat{\mu} = \frac{1}{N} \sum_{n=1}^{N} x_n
$$

• By minimizing the negative log-likelihood, we found: Similarly, we can derive:

sample mean sample variance

- $\hat{\theta} = (\hat{\mu}, \hat{\sigma})$ is the Maximum Likelihood estimate for the parameters of a Gaussian distribution.
	- This is a very important result.
	- Unfortunately, it is wrong...

- To be precise, the result is not wrong, but biased.
- Assume the samples x_1, x_2, \ldots, x_N come from a true Gaussian distribution with mean μ and variance σ^2
	- It can be shown that the expected estimates are then

$$
\mathbb{E}[\mu_{\mathrm{ML}}] = \mu
$$

$$
\mathbb{E}[\sigma_{\mathrm{ML}}^2] = \left(\frac{N-1}{N}\right)\sigma^2
$$

- *The ML estimate will underestimate the true variance!*
- We can correct for this bias:

$$
\hat{\sigma}^2 = \frac{N}{N-1} \sigma_{\rm ML}^2 = \frac{1}{N-1} \sum_{n=1}^{N} (x_n - \hat{\mu})^2
$$

- Maximum Likelihood has several significant limitations.
	- It systematically underestimates the variance of the distribution!
	- E.g., consider the estimate for a single sample:

- We say ML overfits to the observed data.
- *We will still often use Maximum Likelihood, but it is important to know about this effect.*

Probability Density Estimation

- 1. Probability Distributions
- 2. Parametric Methods
- **3. Nonparametric Methods**
	- **a) Histograms**
	- b) Kernel Methods & k-Nearest Neighbors
- 4. Mixture Models
- 5. Bayes Classifier
- 6. K-NN Classifier

Histograms

- Partition the data space into M distinct bins with widths Δ_i and count the number of observations n_i in each bin.
- Then, $p_i = \frac{n_i}{N\Delta_i}$.
- Often the same width is used for all bins.
- This can be done, in principle, for any dimensionality D .

…but the required number of bins grows exponentially with $D!$

The bin width Δ acts as a smoothing factor.

Advantages Limitations

- Very general method. In the limit $(N \to \infty)$, every probability density can be represented.
- No need to store the data points once histogram is computed.

- Rather brute-force.
- Discontinuities at bin edges.
- Choosing right bin size is hard.
- Unsuitable for high-dimensional feature spaces.

Probability Density Estimation

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Kernel Methods and k-Nearest Neighbors

- Data point x comes from pdf $p(\mathbf{x})$.
	- Probability that x falls into small region \mathcal{R} :

$$
P = \int_{\mathcal{R}} p(y) dy \approx p(\mathbf{x}) V
$$

- Estimate $p(\mathbf{x})$ from samples
	- Let K be the number of samples that fall into \mathcal{R} .
	- If the number of samples N is sufficiently large, we can estimate P as:

$$
P = \frac{K}{N} \quad \Rightarrow \quad p(\mathbf{x}) \approx \frac{K}{NV}
$$

 $V:$ volume of \mathcal{R} . For sufficiently small \mathcal{R} , $p(\mathbf{x})$ is roughly constant.

Kernel Methods

• Hypercube of dimension D with edge length h :

$$
k(\mathbf{u}) = \begin{cases} 1, & \text{if } |u_i| \le \frac{1}{2}h, \ i = 1, \dots, D \\ 0, & \text{otherwise} \end{cases}
$$
\n
$$
K = \sum_{n=1}^{N} k(\mathbf{x} - \mathbf{x}_n) \qquad V = \int k(\mathbf{u}) d\mathbf{u} = h^D
$$

• Probability density estimate:

$$
p(\mathbf{x}) \approx \frac{K}{NV} = \frac{1}{Nh^D} \sum_{n=1}^{N} k(\mathbf{x} - \mathbf{x}_n)
$$

• This method is known as Parzen Window estimation.

Nonparametric Methods | Kernel Methods

• In general, we can use any kernel such that

$$
k(\mathbf{u}) \ge 0, \quad \int k(\mathbf{u}) d\mathbf{u} = 1
$$

$$
K = \sum_{n=1}^{N} k(\mathbf{x} - \mathbf{x}_n)
$$

Then, we get the probability density estimate

$$
p(\mathbf{x}) \approx \frac{K}{NV} = \frac{1}{N} \sum_{n=1}^{N} k(\mathbf{x} - \mathbf{x}_n)
$$

• This is known as Kernel Density Estimation.

E.g., a Gaussian kernel for smoother boundaries.

k-Nearest Neighbors

- Fix K , estimate V from the data.
- Consider a hypersphere centered on x and let it grow to a volume V^* that includes K of the given N data points.
- Then

$$
p(\mathbf{x}) \approx \frac{K}{NV^*}
$$

- Side note:
	- Strictly speaking, the model produced by k-NN is not a true density model, because the integral over all space diverges.
	- E.g. consider $K=1$ and a sample exactly on a data point.

$$
V^* = 0 \qquad \Rightarrow \qquad p(\mathbf{x}) \approx \frac{K}{N \cdot 0}
$$

Advantages Limitations

- Very general. In the limit $(N \to \infty)$, every probability density can be represented.
- No computation during training phase
	- Just need to store training set

- Requires storing and computing with the entire dataset.
	- Computational costs linear in the number of data points.
	- Can be improved through efficient storage structures (at the cost of some computation during training).
- Choosing the kernel size/ K is a hyperparameter optimization problem.

Bias-Variance Tradeoff

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References and Further Reading

- More information in Bishop's book
	- Gaussian distribution and ML: Ch. 1.2.4 and 2.3.1-2.3.4.
	- Nonparametric methods: Ch. 2.5.
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